Universal distribution of the number of minima for random walks and Lévy flights

Grégory Schehr Laboratoire de Physique Théorique et Hautes Energies CNRS/Sorbonne Université, Paris

> Journées de Physique Statistique Paris, January 30-31, 2025

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- Anupam Kundu (ICTS-TIFR, Bengaluru)
- Satya N. Majumdar (LPTMS, Univ. Paris-Saclay)

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A. Kundu, S. N. Majumdar, G. S., J. Phys. A: Math. Theor. 58, 035002 (2025)

Outline

Motivations and background

Main results

Sketch of the derivation

Conclusion and perspectives

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Conclusion and perspectives

Study of random landscapes



from V. Ros

Study of random landscapes



Q: what is the number of stationary points ?



Q: what is the number of stationary points ?

Natural and important question in various contexts



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- Natural and important question in various contexts
 - Energy landscapes of complex/disordered systems (e.g., spin-glasses)

Ben Arous, Biroli, Fyodorov, Lacroix-A-Chez-Toine, Le Doussal, Ros, ...



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 - String theory (estimating possible vacua)

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... also relevant for one-dimensional landscapes !

Here: one-dimensional random landscape generated by a random walk

$$x_0 = 0$$
 , $x_n = x_{n-1} + \eta_n$, $n \ge 1$

IID rand. var. distributed with $\phi(\eta),$ continuous and symmetric



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Discrete version of the Sinai model (diffusion in an energy landscape generated by a Brownian motion)



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Mapping to fluctuating interfaces

e.g.,
$$\phi(\eta) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{\eta^2}{2\sigma^2}}$$
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$$P_{\text{joint}}(x_0, x_1, \dots, x_N) = \frac{1}{Z_N} e^{-\frac{1}{2\sigma^2} \sum_{i=0}^{N-1} (x_{i+1} - x_i)^2}$$



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Gibbs measure of an elastic interface (« solid on solid » model)



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JULY 2000

Extremal-point densities of interface fluctuations

Z. Toroczkai,^{1,4} G. Korniss,³ S. Das Sarma,¹ and R. K. P. Zia² ¹Department of Physics, University of Maryland, College Park, Maryland 20742-4111 ²Department of Physics, Virginia Polytechnic Institute and State University, Blacksburg, Virginia 24061-0435 ³Supercomputer Computations Research Institute, Florida State University, Tallahassee, Florida 32306-4130 ⁴Center for Nonlinear Science and Theoretical Division, Los Alamos National Laboratory, Los Alamos, New Mexico 87544



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results only for the average number of minima, but nothing on the fluctuations !

These questions have been widely studied for lattice paths (or equivalently Bernoulli random walks) using combinatorial approaches

Narayana, Krattenthaler, Stanley, ...

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- A different terminology is often used in this context:
 - « turns » stand for « stationary points »
 - « peaks » stand for « local maxima »

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More recent works considered constrained lattice paths

Asymptotics of Bernoulli random walks, bridges, excursions and meanders with a given number of peaks

> Jean-Maxime Labarbe Université de Versailles

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 $^{\blacksquare}$ Exact results for finite number of steps N + central limit theorem in the limit $N \rightarrow \infty$

 $\mathbb{E}(\underline{m}_{N}) = N/4 + O(1)$, $Var(\underline{m}_{N}) = N/16 + O(1)$

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This talk: what about more general constrained random walks with continuous jump/increment distribution ?

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 \blacksquare Stat. of the number of minima $\mathsf{m}_N^{\mathrm{rw}}$ for free random walks of N steps

$$Q^{\text{rw}}(m,N) = \text{Prob}.\,(\mathsf{m}_N^{\text{rw}} = m) = \begin{cases} \frac{1}{2^N} \binom{N+1}{2m+1}, & \text{for } 0 \le m \le \frac{N}{2} \\ 0, & \text{for } m > \frac{N}{2} \end{cases}$$



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(Obviously) universal, i.e., independent of the increment/ jump distribution $\phi(\eta)$



• Stat. of the number of minima m_N^{me} for RW meander of N steps

 $Q^{\text{me}}(m, N) = \text{Prob.}(m_N^{\text{me}} = m \& x_n \ge 0 \text{ for all } 1 \le n \le N)$


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... Generalization of Sparre Andersen theorem

A reminder on Sparre Andersen theorem



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Sparre Andersen theorem (1954) for the « survival » probability

$$q_{N} = \operatorname{Prob.} (x_{1} \ge 0, x_{2} \ge 0, \cdots, x_{N} \ge 0) = \frac{1}{2^{2N}} {\binom{2N}{N}}$$

Universal, i.e., independent of the increment/jump distribution $\phi(\eta)$

Ph. Mounaix, S. N. Majumdar, G. S., J. Phys. A (2020)



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[•] Consider the time of the minimum t_{\min}

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$$t_{\min} = m'' \iff x_m = x_{\min} = \min\{x_0, x_1, \dots, x_n\}''$$

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Probability distribution of the minimum t_{min}

$$P_n(m) = \text{Prob}.(t_{\min} = m) = q(m)q(n - m) , \quad 0 \le m \le n$$

survival proba. up to step n-m

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Taking the generating function w.r.t. n

$$\tilde{q}(z) = \sum_{m \ge 0} z^n q(n)$$

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 $m \ge 0$



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generating function w.r.t. n $\tilde{q}(z) = \sum z^n q(n)$

Taking the generating function w.r.t. n

$$\tilde{q}(z)^2 = \frac{1}{1-z} \Longrightarrow \tilde{q}(z) = \frac{1}{\sqrt{1-z}}$$

•—•local minima (N-step first-passage walk) x_n λ 0 N n-

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rand. var. distributed with $\phi(\eta)$, continuous and symmetric

• Joint proba. of the number of minima m_N^{fp} and the first-passage time

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$$Q^{\text{fp}}(m, N) = \text{Prob}. (\mathsf{m}_N^{\text{fp}} = m \& x_n \ge 0 \text{ for all } 0 \le n \le N - 1 \& x_N < 0)$$

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$$\int \frac{(N-1)!}{2^{2m+N+1} m! (m+1)! (N-2m-2)!} \quad \text{for} \quad 0 \le m \le \frac{N}{2} - 1,$$

$$\begin{bmatrix} m, n \end{pmatrix} = \begin{bmatrix} 0 & \text{for } m > \frac{N}{2} - 1 \end{bmatrix}$$

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Again universal, i.e., independent of the increment/jump distribution $\phi(\eta)$!



$$Q^{\text{fp}}(m) = \sum_{N \ge 2(m+1)} \text{Prob.} (\mathsf{m}_N^{\text{fp}} = m \& x_n \ge 0 \text{ for all } 0 \le n \le N - 1 \& x_N < 0)$$

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Remark:
$$Q^{\text{fp}}(0) = \frac{3}{4}$$

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$$Q^{\rm fp}(m) = \frac{1}{2^{2m+2}} \frac{(2m)!}{m!(m+1)!} \sim \frac{1}{m \to \infty} \frac{1}{4\sqrt{\pi}} \frac{1}{m^{3/2}} \qquad m \ge 1$$

Remark:
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More results for the free random walk

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 - I Joint probability distribution of the nber of minima m_N^{rw} and the nber of maxima M_N^{rw} for a free RW of N steps

$$\sum_{N \ge 2} z^N \sum_{m \ge 0} u^m \sum_{M \ge 0} v^M \operatorname{Prob} . \left(\mathsf{m}_N^{\operatorname{rw}} = m, \mathsf{M}_N^{\operatorname{rw}} = M\right) = z^2 \frac{(2-z) + (u+v) + z \, u \, v}{(2-z)^2 - u v z^2}$$

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Correlation function: $\mathbb{E}(\mathsf{m}_n\mathsf{M}_N) - \mathbb{E}(\mathsf{m}_N)\mathbb{E}(\mathsf{M}_N) = \frac{N-3}{16}$

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• Define $Q_{\pm}^{\mathrm{rw}}(m,N) = \operatorname{Prob} \left(\mathsf{m}_{N}^{\mathrm{rw}} = m \quad \& \quad \operatorname{sign}(\eta_{1}) = \pm \right)$

• Write coupled backward equations for $Q^{\mathrm{rw}}_{\pm}\equiv Q$, for $m\geq 1$, $N\geq 3$

 ${}^{\blacksquare}$ Write coupled backward equations for $Q_{\pm}^{\mathrm{rw}}\equiv Q$, for $m\geq 1$, $N\geq 3$

$$\begin{cases} Q_{+}(m,N) &= \frac{1}{2}Q_{+}(m,N-1) + \frac{1}{2}Q_{-}(m,N-1) \\ Q_{-}(m,N) &= \frac{1}{2}Q_{+}(m-1,N-1) + \frac{1}{2}Q_{-}(m,N-1) \end{cases}$$

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starting from $Q_{+}(m,2) = \delta_{m,0}/2$, $Q_{-}(m,2) = (\delta_{m,0} + \delta_{m,1})/4$

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starting from $Q_+(m,2)=\delta_{m,0}/2$, $Q_-(m,2)=(\delta_{m,0}+\delta_{m,1})/4$

For the free random walk, these recursion relations can be easily solved via generating function techniques

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starting from $Q_{+}(m,2) = \delta_{m,0}/2$, $Q_{-}(m,2) = (\delta_{m,0} + \delta_{m,1})/4$

For the free random walk, these recursion relations can be easily solved via generating function techniques

• Unfortunately, they are much harder to solve for constrained random walks, like meanders (except for some special jump distributions, e.g., $\phi(\eta)=e^{-|\eta|}/2$)

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another method is needed !

A second approach via an auxiliary random walk

The RW built from the local minima

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Note that the number of steps of this effective RW is not fixed

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The effective jump distribution $\psi(y' - y)$ is symmetric and continuous A. Kundu, S. N. Majumdar, G. S. `2024

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Cumulative distribution $Q_{>}^{\text{fp}}(m)$: proba. of observing at least m minima

up to the first-passage time

$$Q^{\rm fp}_{>}(m) = \sum_{m' \ge m} Q_{\rm fp}(m')$$

A. Kundu, S. N. Majumdar, G. S. `2024 Cumulative distribution $Q_{>}^{\rm fp}(m)$: proba. of observing at least m minima up to the first-passage time



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Cumulative distribution $Q_{>}^{\text{fp}}(m)$: proba. of observing at least m minima up to the first-passage time



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Cumulative distribution $Q_{>}^{\rm fp}(m)$: proba. of observing at least m minima up to the first-passage time



A. Kundu, S. N. Majumdar, G. S. `2024 Cumulative distribution $Q_{>}^{\rm fp}(m)$: proba. of observing at least m minima up to the first-passage time $Q_{>}^{\rm fp}(m) = \sum Q_{\rm fp}(m')$

 $m' \ge m$ x_n y_m y_4 y_2 y_3 y_1 \vec{n} $Q_{>}^{\text{fp}}(m) = q_m \times \frac{1}{2}$ $q_m = \text{Prob}. (x_1 \ge 0, x_2 \ge 0, \dots, x_m \ge 0) = \frac{1}{2^{2m}} \binom{2m}{m}$ 10 1 1

$$Q^{\text{fp}}(m) = Q^{\text{fp}}(m+1) - Q^{\text{fp}}(m) = \frac{1}{2^{2m+2}} \frac{(2m)!}{m!(m+1)!}$$

Outline

Motivations and background

Main results

Sketch of the derivation

Conclusion and perspectives

Conclusion and some open questions

Universal results for the statistics of the number of minima for constrained RWs

Our results for the meanders can be seen as an extension of the Sparre Andersen theorem

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Can it be extended to other processes ?

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Our results for the meanders can be seen as an extension of the Sparre Andersen theorem

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Are there some natural extensions of these questions in higher dimensions? see, e.g., extensions of the Sparre Andersen theorem in higher dimensions by Z. Kabluchko, V. Vysostsky, D. Zaporozhets Thank You !